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A NOTE ABOUT THE STRONG CONVERGENCE OF THE  
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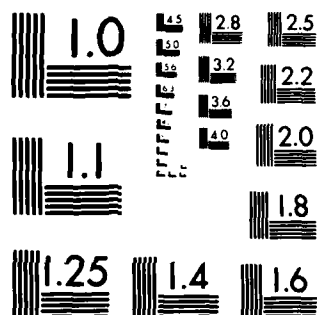
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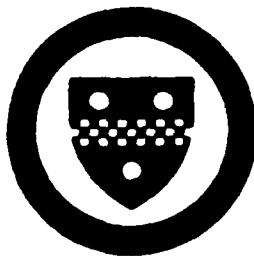
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# A NOTE ABOUT THE STRONG CONVERGENCE OF THE NONPARAMETRIC ESTIMATION OF A REGRESSION FUNCTION

by

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## ABSTRACT

Consider the regression model  $Y_i^* = g(x_i^*) + e_i^*$   $i = 1, 2, \dots, n$ .  $X_i^*$ 's are unordered design variables,  $g$  unknown function defined on  $[0,1]$ .  $\{e_i^*\}$  i.i.d r.v with mean 0 and finite moment of order  $p > 1$ . The asymptotic behavior of estimator  $g_n$  are studied.

Key Words: Nonparametric regression, Kernel estimation, large sample property.

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MATTHEW J. KOPPEL

Chief, Technical Information Division

We consider the regression model

$$Y_i^* = g(x_i^*) + e_i^* \quad i = 1, 2, \dots, n \quad (1)$$

where  $x_i^* \in [0,1]$  are unordered design variables.  $e_i^*$  are iid random variables with zero mean.  $g(x)$  is an unknown function defined on the interval  $[0,1]$ . The problem is to estimate  $g(x)$  through observable variables  $Y_i^*$ .

Denote ordered design variables by  $x_1 \leq x_2 \leq \dots \leq x_n$ , and define  $Y_i = Y_j^*$  and  $e_i = e_j^*$  if  $x_i = x_j^*$  and  $\delta_n = \max_{1 \leq i \leq n+1} (x_i - x_{i-1})$  where  $X_0 \equiv 0$ ,  $X_{n+1} \equiv 1$ . An estimator proposed by Lin and Cheng [1] is given by

$$g_n(x) = \sum_{i=1}^n Y_i \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \quad (2)$$

where  $\{a_n\}$  is a sequence of positive constants converging to 0, as  $n \rightarrow \infty$ ,  $k(z)$  is a kernel density satisfying:

- (i)  $k(z) \geq 0$ ;
- (ii)  $k(z) = 0$  for  $z \notin [-L, L]$  for some positive constant  $L$ ;
- (iii)  $\int_{-L}^L k(z) dz = 1$ .

In fact, it is obtained by the natural estimator

$$h(x) = \begin{cases} Y_1 & x \leq x_1 \\ Y_{i+1} & x_i < x \leq x_{i+1} \\ Y_n & x > x_n \end{cases} \quad i = 1, \dots, n-1 \quad (3)$$

after smoothing with weight function  $a_n^{-1} k[(x-z)/a_n]$ . Lin and Cheng [1] discussed the strong convergence of  $g_n$  and the rates of uniform convergence under a condition



$E|e^*|^p < \infty$ ,  $p \geq 2$  among other conditions. In this note we relax the restrictions on the moments of  $e_i^*$  and improve the results of [1] and [2]. The results are given in the following theorem:

**Theorem** Assume that  $k(z)$  is bounded,  $E|e^*|^p < \infty$  for  $p > 1$ ,  $\delta_n = O(n^{-1})$ ,  $a_n = n^{-d}$ , for  $0 < d < 1 - \frac{1}{p}$ . If  $g(x)$  is continuous in  $[0,1]$ , then for all  $x \in (0,1)$

$$g_n(x) \rightarrow g(x) \quad \text{a.s.} \quad (4)$$

Proof: Denote  $e_i = e_i I(|e_i| \leq i^{1/p})$ .

$\tilde{g}_n = \sum_{i=1}^n \tilde{e}_i \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz$ , where  $I(\cdot)$  is an indicator function. Then

$$\begin{aligned} g_n(x) - g(x) &= g_n(x) - \tilde{g}_n(x) + \tilde{g}_n(x) - E\tilde{g}_n(x) + E\tilde{g}_n(x) - Eg_n(x) + Eg_n(x) - g(x) \\ &\stackrel{\Delta}{=} I_1 + I_2 + I_3 + I_4. \end{aligned}$$

In the following we prove  $I_i \xrightarrow{\text{a.s.}} 0$ ,  $i = 1, 2$  and  $I_j \rightarrow 0$ ,  $j = 3, 4$  separately. First, when  $n$  is large enough

$$\begin{aligned} I_4 &= Eg_n - g = \sum_{i=1}^n g(x_i) \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz - g(x) \\ &= \sum_{i=1}^n \int_{x_{i-1}}^{x_i} [g(x_i) - g(z)] a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \\ &\quad + \sum_{i=1}^n \int_{x_{i-1}}^{x_i} [g(z) - g(x)] a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \end{aligned} \quad (6)$$

By the uniform continuity of  $g(x)$  in  $[0,1]$  and the property of  $k(z)$ , the first term above is less than  $\varepsilon \int_0^1 \frac{1}{a_n} k\left(\frac{x-z}{a_n}\right) dz$ , where  $\varepsilon$  is arbitrary positive constant. The second term in (6) is not greater than

$$\sup_{|z-x|<\delta} |g(z) - g(x)| + \frac{1}{\delta} \sup_{|t|>\frac{\delta}{a_n}} |tk(t)| \int_0^1 |g(z)| dz \\ + \sup_x |g(x)| \int_{|u|>\frac{\delta}{a_n}} k(u) du.$$

So it is easy to see when  $n \rightarrow \infty$ ,

$$I_4 \rightarrow 0 \quad (7)$$

by Holder and Chebyshev inequality.

$$|I_3| \leq \sum_{i=1}^n E\{|e_i| I(|e_i| > i^{1/p})\} \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \\ \leq a_n^{-1} \delta_n \|k\| \sum_{i=1}^n [P(|e_i| > i^{1/p})]^{1-\frac{1}{p}} (E|e_i|^p)^{\frac{1}{p}} \\ \leq a_n^{-1} \delta_n \|k\| \sum_{i=1}^n \frac{1}{i^{1-\frac{1}{p}}} (E|e_i|^p)^{1-\frac{1}{p}} (E|e_i|^p)^{\frac{1}{p}} \\ \leq c a_n^{-1} \delta_n \sum_{i=1}^n \frac{1}{i^{1-\frac{1}{p}}} \leq c a_n^{-1} \delta_n n^{\frac{1}{p}} \rightarrow 0, \quad n \rightarrow \infty \quad (8)$$

where  $\|k\| = \sup_x |k(x)|$ ,  $c$  a constant. Notice that if  $1 < b \leq 2$ , we have  $e^z \leq 1 + z + |z|^b$  when  $z \leq 1$ , so if r.v  $Z_i \leq 1$ ,  $EZ_i = 0$  then  $E(\exp\{Z_i\}) \leq \exp(E|Z_i|^b)$ . Now, take  $b < p$ ,  $d_n = \ell_n n^{\frac{1}{p}} \ell_n$ ,  $Z_i = d_n(\tilde{e}_i - E\tilde{e}_i) \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz$ . By the choice of  $d_n$  and the conditions on  $\delta_n$  and  $a_n$  in the theorem we know that  $Z_i \leq 1$  when  $n$  is large enough. Thus, we have

$$E\left(\prod_{i=1}^n \exp\{Z_i\}\right) = \prod_{i=1}^n E(\exp\{Z_i\}) \leq \prod_{i=1}^n \exp(E|Z_i|^b) \\ \leq \prod_{i=1}^n \exp\{d_n(d_n \|k\| \delta_n a_n^{-1})^{b-1} E|e_i|^b \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz\} \\ \leq \exp\{c d_n(d_n \delta_n a_n^{-1})^{b-1}\}$$

For given  $\varepsilon > 0$ , when  $n$  is large enough,

$$\begin{aligned}
 P(I_2 > \epsilon) &\leq e^{-d_n \epsilon} E(\exp\{d_n I_2\}) \\
 &= e^{-d_n \epsilon} E\left(\prod_{i=1}^n \exp\{Z_i\}\right) < e^{-\frac{1}{2} d_n \epsilon}.
 \end{aligned}$$

So, we have

$$\sum_n P(I_2 > \epsilon) < \sum_n e^{-\frac{1}{2} d_n \epsilon} < \infty$$

by Borel-Cantelli lemma

$$\limsup_{n \rightarrow \infty} I_2 \leq 0 \quad \text{a.s.}$$

Similarly, it is true that

$$\liminf_{n \rightarrow \infty} I_2 \geq 0 \quad \text{a.s.}$$

Therefore

$$I_2 \rightarrow 0 \quad \text{a.s.} \quad (9)$$

Since  $E|e_i|^p < \infty$ , we know that  $\sum_n P(|e_i| > j^{\frac{1}{p}}) < \infty$ . By Borel-Cantelli lemma we have  $P(|e_i| > i^{\frac{1}{p}}, \text{ i.o.}) = 0$ . Hence

$$\sum_{i=1}^{\infty} e_i^2 I(|e_i| > i^{\frac{1}{p}}) < \infty \quad \text{a.s.} \quad (10)$$

Finally,

$$|I_1|^2 = \left| \sum_{i=1}^n (e_i - e_i) \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \right|^2.$$

By Schwartz inequality,

$$|I_1|^2 \leq \sum_{i=1}^n \left( \int_{x_{i-1}}^{x_i} a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \right)^2 \sum_{i=1}^n e_i^2 I(|e_i| > i^{\frac{1}{p}})$$

$$\begin{aligned}
&\leq a_n^{-1} \delta_n \|k\| \int_0^1 a_n^{-1} k\left(\frac{x-z}{a_n}\right) dz \sum_{i=1}^{\infty} e_i^2 I(|e_i| > i^{\frac{1}{p}}) \\
&\leq c a_n^{-1} \delta_n \sum_i e_i^2 I(|e_i| > i^{\frac{1}{p}}).
\end{aligned}$$

By (10) we have

$$|I_1|^2 \rightarrow 0 \text{ a.s.} \quad (11)$$

From (5), (7), (8), (9), and (11), the theorem follows.

## REFERENCES

- [1] CHENG, Kuang-Fu and LIN, Pi-Erh, Z. Wahrscheinlichkeitstheorie verw. Gebiete, 57, 223-233 (1981).
- [2] BENEDETTI, J. Roy. Statist. Soc. B39, 243-253, (1977).

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